

Initial Development and Reliability of a Motivation for Weight Loss Scale

Andrea H. Meyer^{a,b} Simone Weissen-Schelling^b Simone Munsch^c Jürgen Margraf^b

^a Division of Applied Statistics in Life Sciences,

^b Department of Clinical Psychology and Psychotherapy, Faculty of Psychology, University of Basel,

^c Institute of Psychology, Clinical Child and Adolescent Psychology, University of Lausanne, Switzerland

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Summary

Objective: We aimed at developing and evaluating a questionnaire assessing health and appearance as the two main reasons for weight loss in overweight and obese individuals. **Methods:** Using data from two representative telephone surveys in Switzerland, the factorial structure of this questionnaire was analyzed by exploratory and confirmatory factor analysis. The model obtained was cross-validated with data from a second representative Swiss survey and multigroup analyses according to sex, age, BMI and regional language subgroups were performed. **Results:** This led to a 24-item, 3-factor solution, with factors labeled 'health', 'appearance in relation to others', and 'appearance in relation to oneself'. Internal consistency and test-retest reliability were good. **Conclusions:** To the best of our knowledge, this is the first validated questionnaire assessing overweight and obese individuals' reasons for weight loss. It should be further tested whether using this questionnaire as a pretreatment assessment device will help in tailoring treatments to individuals, thereby increasing treatment adherence and success.

Introduction

In the past few decades, overweight and obesity have reached epidemic proportions worldwide [1]. As both are associated with increased morbidity in the areas of physical and mental

health [2–4], health care costs for weight-related illnesses are soaring [5]. Moreover, overweight and obesity are associated with decreased quality of life [6–7], lower socioeconomic status [8] and stigmatization [9], leading to an economic as well as psychological burden for societies and individuals all over the world.

Previous research has shown that weight loss is an important concern for many of the individuals affected by overweight and obesity, with women trying to lose weight at a lower BMI than men [10–12]. Despite the fact that many individuals want to lose weight and many treatment options are available, weight loss treatments are only of moderate success in the short and long term [13–15], and dropout rates are usually high [16–17]. Therefore, it is important to understand the factors contributing to treatment adherence and success. Guidelines [18] suggest that reasons and motivation for weight loss represent meaningful characteristics of patients that should be assessed prior to any weight loss intervention in order to identify those individuals who are really ready – that is, motivated – for weight loss and those who are more ambiguous and at risk of dropping out of treatment.

Data from weight loss intervention studies and telephone surveys indicate that the perception of obesity as a health risk [19, 12] and dissatisfaction with own appearance [20–23] are the two main motivating factors for weight loss attempts in about 50–85% and 15–36% of overweight and obese individuals, respectively [21–23]. Individuals citing either health or appearance as their number one reason for weight loss differ in regard to self-esteem [23–24], body dissatisfaction [22–24], and their BMIs [23]. Moreover, motivating factors for weight loss seem to differ with age [19, 24]. A qualitative report covering people's reasons for entering a weight loss regime [25] found that those who were most successful in losing a significant amount of weight were most likely to indicate health as reason for weight loss.

Thus, despite the urgent need to enhance treatment outcomes in overweight and obese patients and the often-cited impact of an individual's motivation for weight loss on treatment success, there is a lack of reliable and valid assessment devices in this area of research. The best-known measure assessing a person's willingness and preparedness to engage in the behavioral practices required to lose weight is the Dieting Readiness Test (DRT) [26]. However, this scale assesses readiness for weight loss on different dimensions but not a person's reasons for losing weight. Moreover, although research studies agree that health and appearance are the two main reasons for weight loss, there is no indication of how they are correlated.

The aim of the present study therefore was to develop and evaluate a questionnaire that considers health and appearance as different aspects of motivation for weight loss. By using this questionnaire as a pretreatment assessment device, a person's motives for weight loss can be identified, and a weight loss program's components could be tailored to the individual, thereby enhancing treatment success and reducing the usually high dropout rates.

Participants and Methods

Participants

In 2004, two telephone surveys were conducted to obtain representative samples from the German- and French-speaking parts of Switzerland. In both surveys, participants were selected according to the random-quota method, with age, sex, geographical region, and size of community being the quota characteristics. Participants were selected from the database of current landline telephone accounts. To be included in the survey, participants had to be between 15 and 74 years of age. Participation was voluntary. In the first survey, conducted in January 2004, 1,000 participants were interviewed. In the second survey, conducted in June 2004, the corresponding sample size was 800. For our analysis we used the interviewees reporting a BMI of ≥ 25 kg/m², that is 355 (36%) from survey 1 and 232 (29%) from survey 2. The interviewees of survey 2 were re-interviewed 1 week later, where 140 of them participated again, corresponding to a response rate of 60.3%. Sample characteristics are displayed in table 1.

Procedure

Both telephone surveys were conducted by IHA GfK AG, a professional research institute in Switzerland. In survey 1, the questionnaire consisted of 39 questions concerning height, weight, reasons for wanting to lose weight, dieting history, binge-eating episodes, physical activity and several sociodemographic characteristics. In the first wave of the second survey questions concerning height, weight and reasons for wanting to lose weight as well as several sociodemographic questions were asked, whereas in the second wave only the questions concerning reasons for wanting to lose weight were presented. For the construction and validation of our questionnaire only questions concerning reasons for wanting to lose weight were analyzed. These questions were exactly the same in both surveys and in both waves of survey 2 with respect to the number asked and the phrasing.

Scale Construction

We used a deductive scale development strategy based on theoretical considerations. After a thorough review of existing literature regarding

Table 1. Sample characteristics of overweight and obese interviewees of surveys 1 and 2

Variable	Survey 1 (N = 355)	Survey 2, first wave (N = 232)	Survey 2, second wave (N = 140)
Gender, % female	35.8	38.4	42.1
Language region, % German speaking	73.2	73.7	81.4
Mean age, years (SD)	48.3 (14.8)	49.5 (14.6)	50.1 (14.1)
Mean BMI, kg/m ² (SD)	28.3 (3.8)	29.0 (3.6)	28.7 (2.8)
Employment status, %			
Full time	47.9	43.5	43.6
Part time	16.9	16.4	18.6
Unemployed ^a	35.2	40.1	37.9
Educational attainment, %			
Primary or lower	16.3	13.8	14.3
Secondary/upper	70.7	66.0	67.1
Secondary/tertiary	12.9	20.2	18.5
Living situation, %			
Family	49.4	45.3	45.0
Couple	36.4	41.4	41.4
Single	14.4	13.4	13.6

^aMostly retired persons, housewives and students.

obese individuals' motivational reasons for weight loss [19, 22–25], we generated a wide array of possible reasons for weight loss that could be allocated to either the health or appearance motive. This pool of 39 items was then presented to the interviewees (each item consisting of the statement 'I want to lose weight ...', followed by a different weight loss reason). All items had to be answered on a four-point Likert scale with the values 1: 'absolutely not', 2: 'somewhat', 3: 'moderately' and 4: 'strongly', indicating how much participants identified with the different statements.

Statistical Analysis

Data Preparation

Of the 355 overweight and obese participants captured in survey 1, 25 datasets contained at least one missing answer and were eliminated, leaving a total of 330 observations for subsequent analysis. For the same reason we eliminated 17 participants from the first wave of survey 2, leaving a total of 215 observations for analysis. Data preparation for both surveys included the log transformation of the 39 items to achieve normality and homoscedasticity and the subsequent elimination of 10 (survey 1) and 6 (first wave of survey 2) observations that were either univariate or multivariate outliers [27]. Thus, the sample sizes for the analysis of survey 1 and the first wave of survey 2 were 320 and 209, respectively.

Construction of Final Questionnaire and Evaluation of Scale

First, exploratory factor analysis (EFA) based on the principal axes factoring method with varimax rotation was used to extract the factorial structure of the questionnaire items [28]. EFA was done using data from survey 1, and the factorial structure obtained was analyzed using the Kaiser-Guttman criterion, the scree plot and parallel analysis [29] for identification of the number of factors needed. The obtained model was subsequently tested for goodness of fit using confirmatory factor analysis (CFA) and then cross-validated using the data from the first wave of survey 2. No modifications were made to the model to improve goodness of fit.

Table 2. Factor loadings of the three-factor solution regarding reasons for wanting to lose weight

Item no.	Item label	Factor loadings		
		I	II	III
<i>Factor I: health</i>				
3.	Because it is commonly said that being overweight is unhealthy	0.597	0.153	0.285
5.	To be healthier	0.685	0.134	0.280
9.	To be more agile	0.662	0.175	0.291
10.	For health reasons	0.839	0.084	0.140
15.	Because I read that it is healthier	0.586	0.219	0.107
20.	To decrease my health risks	0.784	0.037	0.11
25.	To live long	0.582	0.230	0.138
<i>Factor II: appearance in relation to others</i>				
1.	Because acquaintances have advised me to	0.127	0.622	0.154
4.	To not attract attention	0.225	0.595	0.274
6.	Because I'll be more successful in my job	0.199	0.626	0.13
7.	So I will be accepted by society	0.177	0.747	0.209
8.	To dare to socialize again	0.072	0.717	0.254
12.	Because I would be luckier in love	0.143	0.631	0.289
16.	To be more appreciated/liked	0.219	0.644	0.246
21.	To have more friends	0.129	0.589	0.159
28.	To have better success with others	0.114	0.681	0.306
39.	So that other people will think better of me	0.042	0.727	0.217
<i>Factor III: appearance in relation to oneself</i>				
2.	To be more attractive	0.288	0.258	0.611
14.	To like to look at myself in the mirror again	0.263	0.202	0.694
23.	Because I want to like myself more	0.130	0.309	0.678
27.	Because I want to be more attractive	0.238	0.285	0.677
30.	To be able to dress more fashionably	0.300	0.237	0.559
31.	To fit into my clothes again	0.194	0.402	0.551
32.	To feel more self-confident	0.288	0.258	0.611

To test for factor invariance with respect to factor loadings and factor variances and covariances between sex, age, obesity and regional language groups, we performed several multigroup analyses within the CFA framework according to Byrne [30], comparing the following subgroups: male versus female (sex), <50 years versus ≥50 years (age), BMI < 30 kg/m² versus BMI ≥ 30 kg/m² (BMI) and German versus French speaking (language region). To perform EFA and CFA, we used the software packages SPSS 14 [31] and AMOS 5 [32], respectively.

Reliability

Internal consistency of items within each factor and of all items combined was computed using Cronbach's alpha [33]. To assess the test-retest reliability (stability of measure over time) of the questionnaire, we compared the data sets from the two waves of the second survey. There were 140 valid cases available for both waves. 23 cases were eliminated because they contained at least one missing answer. Data preparation included the log transformation of the sum of the 24 items for each wave to achieve normality and homoscedasticity, followed by the elimination of another 8 cases that were either univariate or multivariate outliers [27]. The final sample size covering both waves contained 109 observations. We used Pearson's correlation coefficients and intraclass correlation coefficients to assess test-retest reliability for each of the factors obtained as well as for all items combined. In addition, paired t-tests were used to test for changes in mean values between the first and second wave together with repeatability coefficients, defined as 1.96 times the standard deviation of the differences between the values of the first and second wave [34]. Re-

peatability coefficients denote the limits within which 95% of the differences between pairs of measurements are expected.

Results

Exploratory Factor Analysis

We first ran an EFA on all 39 items (results available from the authors on request). Subsequently, 15 items had to be removed because they either had factor loadings less than 0.55 or loaded on more than one factor [35]. Running EFA based on the remaining 24 items led to three factors: 'health', 'appearance in relation to others' and 'appearance in relation to oneself' (table 2), with eigenvalues greater than 1 (values between 5.2 and 3.7) consisting of 7, 10, and 7 items, respectively. Parallel analysis confirmed this result. The three factors accounted for 21.6, 16.2, and 15.3% of the total item variance in the rotated factor solution.

Reliability

Internal consistencies for factors I–III were 0.88, 0.91 and 0.89, and 0.93 if all 24 items were combined. Item-total corre-

Table 3. Fit indices of the unconstrained baseline models of each subgroup

	N	χ^2	df	CFI	TLI	RMSEA	P_{close}	AIC
<i>Full sample</i>								
Survey 1	320	530.5	249	0.930	0.923	0.060	0.013	632.5
Survey 2, first wave, cross-validated	209	571.5	249	0.899	0.888	0.079	0.001	673.5
<i>Subgroups (survey 1)</i>								
<i>Sex</i>								
Male	208	453.2	249	0.913	0.903	0.063	0.012	603.2
Female	112	403.3	249	0.912	0.902	0.075	0.002	553.3
<i>Age</i>								
<50 years	170	398.5	249	0.928	0.920	0.060	0.076	548.5
≥50 years	150	459.0	249	0.898	0.887	0.075	0.000	609.0
<i>BMI</i>								
<30 kg/m ²	255	502.8	249	0.922	0.913	0.063	0.004	652.8
≥30 kg/m ²	65	378.1	249	0.859	0.843	0.090	0.001	528.1
<i>Language region</i>								
German speaking	237	473.6	249	0.924	0.916	0.062	0.012	575.6
French speaking	83	412.2	249	0.868	0.853	0.089	0.000	514.2

CFI = Comparative Fit Index; TLI = Tucker and Lewis Index; RMSEA = Root mean square error of approximation; P_{close} = P for test of close fit; AIC = Akaike Information Criterion.

lations for items within each factor ranged between 0.59 and 0.77 and were thus much higher than 0.30, the value that is considered problematic [36]. Test-retest reliability was high, with correlation coefficients ranging between 0.75 and 0.83 and intraclass correlation coefficients ranging between 0.74 and 0.82. Also, mean differences between the two waves ranged between -0.03 and 0.03 for the three subscales and did not differ from 0 for any of them (paired t-test, $p \geq 0.20$ for all three subscales). Corresponding repeatability coefficients ranged between 0.19 and 0.28 for the three subscales.

Confirmatory Factor Analysis

The three-factor model obtained from EFA was tested using the data of survey 1 and resulted in reasonably good fit indices (table 3, first line). Factor loadings varied between 0.59 and 0.83. Intercorrelations among the three factors were 0.43 (factor I vs. factor II), 0.71 (factor II vs. factor III), and 0.58 (factor I vs. factor III). They were all significantly different from 1, suggesting sufficient discriminant validity to justify three different scales. In addition, the three-factor model fitted the data significantly better than a model consisting of one factor only ($\Delta\chi^2 = 975.9$, $\Delta df = 3$, $p \leq 0.001$, for the difference between the two models).

Cross-Validation

When the model obtained from CFA was fit to the data of the first wave of survey 2, the resulting fit was only slightly worse than when using the original data set from survey 1, suggesting a stable model (table 3, second line). Factor loadings for this model varied between 0.57 and 0.83. Intercorrelations among the three factors were somewhat higher than for sur-

vey 1 with values 0.63 (factor I vs. factor II), 0.78 (factor II vs. factor III), and 0.69 (factor I vs. factor III), but they were still all significantly different from 1. Again, this model fitted the data significantly better than a model with just one factor ($\Delta\chi^2 = 520.7$, $\Delta df = 3$, $p \leq 0.001$). Internal consistencies for the three factors were 0.91, 0.92 and 0.90, and 0.95 if all items were combined.

Factor Invariance of the Three-Factor Model

To test whether the proposed three-factor model was equivalent across subgroups with respect to factor loadings and factor variances and covariances, we first determined goodness-of-fit statistics of the baseline models in each subgroup. Model fits were reasonable for all subgroups except for the subgroups BMI ≥ 30 kg/m² and French-speaking region, where sample sizes were small (table 3).

Goodness-of-fit statistics of the three unconstrained two-group models were all reasonably good (sex: $\chi^2 = 856.9$, $df = 498$, GFI = 0.82, CFI = 0.91, TLI = 0.90, RMSEA = 0.048, $P_{\text{close}} = 0.77$, RMR = 0.016; age: $\chi^2 = 857.6$, $df = 498$, GFI = 0.82, CFI = 0.91, TLI = 0.91, RMSEA = 0.048, $P_{\text{close}} = 0.76$, RMR = 0.016; BMI: $\chi^2 = 883.2$, $df = 498$, GFI = 0.83, CFI = 0.91, TLI = 0.90, RMSEA = 0.049, $P_{\text{close}} = 0.577$, RMR = 0.019; language region: $\chi^2 = 887.3$, $df = 498$, GFI = 0.82, CFI = 0.91, TLI = 0.90, RMSEA = 0.050, $P_{\text{close}} = 0.545$, RMR = 0.017). Constraining all these models to have equal factor loadings did not significantly worsen model fits (table 4). When in addition equal factor variances and covariances between the groups were imposed, model fits again did not significantly worsen for age, BMI and regional language groups. Between males and females, however, we detected invariance with respect to factor

Table 4. Goodness-of-fit statistics for tests of invariance across groups

Model description	χ^2	df	$\Delta\chi^2$	Δdf	Significance
Sex					
Unconstrained, combined model	856.9	498	–	–	–
Factor loadings constrained to be equal	873.0	519	16.1	21	0.76
Factor loadings + variances constrained ^a	880.1	522	23.2	24	0.51
Factor loadings + variances + covariances constrained	898.0	525	41.2	27	0.040
Age					
Unconstrained, combined model	857.6	498	–	–	–
Factor loadings constrained to be equal	884.1	519	26.5	21	0.19
Factor loadings + variances + covariances constrained	892.6	525	35.1	27	0.14
BMI					
Unconstrained, combined model	883.2	498	–	–	–
Factor loadings constrained to be equal	904.3	519	21.1	21	0.45
Factor loadings + variances + covariances constrained	916.9	525	33.7	27	0.18
Language region					
Unconstrained, combined model	887.3	498	–	–	–
Factor loadings constrained to be equal	916.6	519	29.2	21	0.11
Factor loadings + variances + covariances constrained	922.2	525	34.8	27	0.14

$\Delta\chi^2$ = difference in chi-square values, Δdf = difference in degrees of freedom, relative to the unconstrained, combined model.
^aAs the term ‘Factor loadings + variances + covariances constrained’ was significant, we additionally tested for invariant variances alone.

variances and covariances. Separating the effects of equal factor variances from those of equal covariances among factors we found that it was the covariances among the factors, and not the factor variances that differed between males and females (see 3rd and 4th line in table 4). Correlations (i.e. standardized covariances) among factors for males and females were as follows: males: $r = 0.41$ (health vs. appearance in relation to others), $r = 0.79$ (appearance in relation to others vs. in relation to oneself), $r = 0.50$ (health vs. appearance in relation to oneself); females: $r = 0.46$ (health vs. appearance in relation to others), $r = 0.65$ (appearance in relation to others vs. in relation to oneself), $r = 0.74$ (health vs. appearance in relation to oneself).

Discussion

The aim of the present study was to develop and evaluate a questionnaire assessing motivation for weight loss in overweight and obese individuals, as there has been up to now a lack of reliable and validated questionnaires in this area of research. Based on factor analysis of responses to telephone interviews in two representative samples of Swiss inhabitants ($N = 320$ or $N = 209$), we constructed a 24-item questionnaire with items loading on three factors that accounted for 21.6, 16.2 and 15.3% of the total item variance. Results of CFA, internal consistency, test-retest reliability, discriminant validity and cross-validation confirmed the reliability and stability of the questionnaire with its three subscales.

The first factor identified contains items covering health reasons for weight loss. Given that in previous research 50–

85% of obese individuals trying to lose weight indicated health as their primary reason [19, 21–23], health seems indeed to be a strong motivator and its assessment based on a distinct subscale seems justified. Even though interviewees were also presented with items covering aspects of weight loss motivation due to advice from others (‘because my general practitioner advised me to’, ‘because my partner is worried’), these items were not relevant for scale construction. Health thus seems to be a personal issue that might be more influenced by the short- and longer-term consequences of overweight and obesity than by pure advice obtained from others.

The second factor deals with items focusing on the improvement of human relationships due to weight loss and was therefore named ‘appearance in relation to others’. The importance of such a factor is clearly supported by research on quality of life, stigmatization, and career problems of overweight and obese patients [7–9]. By losing weight these individuals hope to handle the weight-related problems they are confronted with in their daily social interactions.

The third factor refers to overweight and obese individuals’ wishes to be more attractive and more likeable to themselves. This factor seems to reflect individuals’ own body dissatisfaction, which they hope to improve by losing weight [37], and was thus labeled ‘appearance in relation to oneself’. Although in both CFAs (survey 1 and first wave of survey 2), the second and third factor were more highly correlated than either of these two factors with the first, the identification of two distinct appearance-related factors seems warranted, as the correlation was less than 0.8 and significantly lower than 1 for both surveys.

Multigroup analyses showed that factor loadings as well as factor variances and covariances did not differ between the subgroups defined for age, BMI and regional language, which justifies the generalizability of the three factors across ages and between non-obese and obese patients as well. These findings should, however, be interpreted with caution as the sample size, especially regarding the subgroups with BMI ≥ 30 kg/m² (n = 65) and the French-speaking region (n = 83), was small, thus impeding the detection of non-invariance. Thus, our questionnaire should be further validated in an enlarged sample of obese patients. Regarding the sex subgroups, results from multigroup analyses rejected invariance among covariances for females and males. Males showed a particularly high correlation between the two factors dealing with appearance as the reason for weight loss. It can be suggested that men might be less susceptible for societal pressure to be thin than woman and thus might focus less on different aspects of appearance. In contrast, females showed a particularly high correlation between the factors health and appearance in relation to oneself as reasons for weight loss.

Several limitations of our study should be borne in mind: First, when we generated the 39 items for the survey, we did not have proved experts in this field who judged the suitability of our items. We were thus not able to assess the content validity of our scales [38]. Second, the sample sizes, especially of survey 2, were rather small for CFA, given the rough guideline of at least 5–10 cases for each parameter to be estimated, even though another rule of thumb states that sample sizes of 100–200 should be the lower limit [39]. Similar arguments can be made for multigroup analyses. Third, we did EFA and

CFA using the same sample from survey 1, but nevertheless cross-validation of the model with data from survey 2 supported the stability of our model. Fourth, the normal- or underweight interviewees of both surveys were not asked the questions regarding reasons for wanting to lose weight. So, our results are restricted to a population of overweight persons for whom weight loss is of primary importance regarding health and social consequences of overweight and obesity. Finally BMI values were based on self-reported information and might be distorted. Thus our prevalence rates of 36% of overweight or obese participants have to be handled with caution, even though the prevalence rates are comparable with prevalences in other European countries.

The reasons and motivating factors for participation in weight loss programs are meaningful characteristics of overweight and obese individuals, and guidelines suggest a pre-treatment assessment [18]. The 24-item questionnaire is the first validated assessment device in this area of research. As a next step it should be used in a clinical survey to assess its validity in treatment settings. In doing so, it could be used to target those individuals who may benefit from additional motivational work before starting a treatment, or to tailor interventions to participants' particular motivational reasons in order to increase adherence rates usually observed in weight loss treatments.

Disclosure

The authors declared no conflict of interest.

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